

Industry and Innovation





ISSN: 1366-2716 (Print) 1469-8390 (Online) Journal homepage: www.tandfonline.com/journals/ciai20

Do China's special economic zones increase incentives to invest in R&D?

Katrin Hussinger & Lorenzo Palladini

To cite this article: Katrin Hussinger & Lorenzo Palladini (21 May 2025): Do China's special economic zones increase incentives to invest in R&D?, Industry and Innovation, DOI: 10.1080/13662716.2025.2503283

To link to this article: https://doi.org/10.1080/13662716.2025.2503283

© 2025 The Author(s). Published by Informa UK Limited, trading as Taylor & Francis Group.



0

Published online: 21 May 2025.

(
	رم

Submit your article to this journal 🖸

Article views: 45



View related articles

View Crossmark data 🗹

RESEARCH ARTICLE

OPEN ACCESS Check for updates

Routledge

Taylor & Francis Group

Do China's special economic zones increase incentives to invest in R&D?

Katrin Hussinger^{a,b,c} and Lorenzo Palladini^d

^aDepartment of Economics and Management, University of Luxembourg (Luxembourg), Luxembourg, Luxembourg; ^bDepartment of Managerial Strategy and Innovation, K.U. Leuven, Leuven, Belgium; ^cEconomics of innovation and industrial dynamics, Centre for European Economic Research (ZEW), Mannheim, Germany; ^dHouse of Innovation, Stockholm School of Economics (Sweden), Stockholm, Sweden

ABSTRACT

China's special economic zones (SEZs) have been established to foster business growth and innovation by improving the institutional context of specific sub-regional areas. We examine the effect of SEZs on the contribution of research and development (R&D) to the market value of firms located in these areas. The market value reflects investors' expectations of future returns to R&D, providing crucial information for strategic investment decisions. Larger R&D contributions to the market value create stronger incentives for firms to invest in innovation. Empirical results suggest that the contribution of R&D to the market value increases through the SEZs program, particularly for R&D intensive firms. This suggests that regional policies, while increasing incentives to innovate, may widen the gap between less and more R&D intensive firms, potentially impacting competition and long-term growth.

KEYWORDS

Special economic zones (SEZs); China; market value; R&D; institutional development; innovation incentives

JEL Classification 032: R58

1. Introduction

Economic development policies targeted at specific geographical regions became popular over the past decades (Farole and Akinci 2011; United Nations Conference on Trade and Development 2019). In particular, since 1979, China has made extensive use of its so-called Special Economic Zones (SEZs) program¹ (Hua, Partridge, and Sun 2023) to tackle the challenges that China-based businesses faced due to the weak institutional environment characterised by poor enforcement of property rights protection, underdeveloped financial systems, and a weak economic infrastructure (Alder, Shao, and Zilibotti 2016). A weak institutional environment can hinder innovation and business growth by creating uncertainties and regulatory burdens, discouraging investment in research and development (R&D), and stifling entrepreneurial activities (Acharya and Subramanian 2009; Acharya, Baghai, and Subramanian 2013; Donges, Meier, and Silva 2023; Moser 2005; North 1991). The SEZs program was established to counteract such institutional inefficiencies so that entrepreneurial and innovation activities could flourish in the selected geographical regions (Alder, Shao, and Zilibotti 2016; Barbieri, Pollio, and Prota 2020; Hartwell 2018; Y. Liu and Jin 2022; Y. Liu, Wu, and Zhu 2023; Lu, Wang, and Zhu 2019; Ma, Cao, and Zhao 2023; Song et al. 2020).

In this paper, we investigate whether the SEZs program achieved one of its objectives of promoting and incentivising innovation within firms through an improvement of the regional institutional context.² Considering that incentives to innovate are mostly intangible and difficult

© 2025 The Author(s). Published by Informa UK Limited, trading as Taylor & Francis Group.

CONTACT Lorenzo Palladini 🙆 lorenzo.palladini@hhs.se 🖃 House of Innovation, Stockholm School of Economics, Norrtullsgatan 2, Stockholm 113 29, Sweden

¹Duranton and Venables (2018), 25, footnote 19) define SEZs broadly as 'any special district with favourable fiscal or institutional treatment.' This definition, which is widely used in the literature (see e.g. also Hua, Partridge, and Sun 2023), includes development zones (DZs). In the literature, SEZs and DZs are typically pooled together. For example, Lu, Wang, and Zhu (2019), 332) refer to SEZs as follows: 'there are several types of SEZs [...]. Economic and technological *development zones* (ETDZs) are broadly defined zones with a wide spectrum of investors. High-tech industrial *development zones* (HIDZs) are intended to promote high-tech industrialization [...]. Specialized *industrial zones* (SIZs) are cluster-type industrial parks aiming to develop particular industries, which should be consistent with local comparative advantages. [...] Export *processing zones* (EPZs) [...] are solely for export processing (to develop export-oriented industries)' (see also footnote 2 in Lu, Wang, and Zhu 2019), or the literature section in; Ma, Cao, and Zhao 2023). All the above-mentioned types of DZs are included in our definition of SEZs. We excluded national SEZs because they affect the entire economy and are implemented with regional heterogeneity which renders causal analysis difficult.

²The SEZs program had also other goals such as increasing foreign and domestic investment, and promoting international trade on top of stimulating technological cooperation and innovation (Lu, Wang, and Zhu 2019).

This is an Open Access article distributed under the terms of the Creative Commons Attribution License (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited. The terms on which this article has been published allow the posting of the Accepted Manuscript in a repository by the author(s) or with their consent.

2 😣 K. HUSSINGER AND L. PALLADINI

to observe,³ we focus on the effect of the SEZs program on the contribution of firms' investment in R&D to its market value. A larger contribution of R&D investment to the firm's market value reflects increased expected future returns and, in turn, incentivises the focal firm to invest more in R&D (Chadha and Oriani 2010; Munari and Oriani 2005). This logic builds on the insight that decision-makers learn from new information through secondary market prices and use this information to guide their future strategic decisions (Bai, Philippon, and Savov 2016; Bond and Goldstein 2015; Bond, Edmans, and Goldstein 2012; Dang and Xu 2018). In this context, as stock market prices are a relevant source of information also for firm insiders,⁴ a larger contribution of R&D investment to the firm's market value represents a strong incentive for managers to adjust their future strategic investment decisions and engage more strongly in R&D.

A larger contribution of R&D to the market value of firms is beneficial for firms as it improves the innovating firm's credibility and reputation (Arora, Fosfuri, and Gambardella 2001; Long 2002), serves as an indication of the firm's ability to convert research investments into potentially valuable knowledge (Levitas and McFadyen 2009), and eases the process of securing funding for future R&D. This, in turn, results in lower capital costs, ultimately contributing to the generation of future profits and firm growth (Hottenrott, Hall, and Czarnitzki 2016; Lev, Sarath, and Sougiannis 2005).

In sum, a large contribution of R&D to the market value of firms generates incentives for firms to invest in R&D. If SEZs increase the contribution of R&D to the market value of firms, they reach the objective of fostering innovation and, therewith, put the region on track to achieve enhanced business activity and economic growth.

This study, therefore, investigates whether regional institutional improvements through SEZs increase the contribution of R&D to the market value of a firm, reflecting strengthened incentives for firms to further invest in R&D. Differently from other studies, we refrain from investigating whether SEZs lead to a direct increase in R&D investment (Wu, Liu, and Huang 2021) because R&D projects need planning and potentially the acquisition of specialised equipment and human capital so that their effect will occur with a delay and will be spread over several years. Given that stock markets adjust quickly to changes in the business environment, the market value is a more direct and immediate measure of the incentives that firms have to innovate resulting from the SEZs program.

The SEZs program facilitates the distinction between 'treated' firms – those located in regions targeted by the program – and 'untreated' firms in non-SEZ regions. This distinction enables the use of causal methods by comparing the contribution of R&D to the market value of treated firms with that of untreated firms (the control group). By doing so, we can abstract from different economic trends over time and isolate the effect of the SEZs. We employ a geographic boundary discontinuity (BD) framework (Holmes 1998; Keele and Titiunik 2015; Lu, Wang, and Zhu 2019) which relies on a control group of untreated firms that are geographically close to regions that were turned into SEZs. The underlying assumption is that economic conditions alter abruptly at the borders of an SEZ due to differences in policies, regulations, or infrastructure, while at the same time, the firms themselves are very similar on both sides. The BD design is powerful in our context because SEZs borders do not coincide with administrative boundaries (Lu, Wang, and Zhu 2019). To demonstrate the robustness of our results across different methods, we employ a coarsened exact matching approach and a random sampling procedure involving 500 randomly selected control samples of three firms for each treated firm. Additionally, we check the robustness of these results by using different geographical bands in which control firms are allowed to be located, by expanding the treatment group to include firms with subsidiaries located in SEZs, and by conducting heterogeneity analyses based on various firm and region characteristics.

³While the outcomes of innovation efforts, such as new products or patents, are observable, incentives generated by policies such as the SEZs program are mostly intangible in nature so that they cannot be easily observed. More specifically, the SEZs program improves the regional institutional environment which should facilitate innovation activities because R&D can be more easily financed and the returns to innovation can be more easily appropriated.

⁴As explained by Bond, Edmans, and Goldstein (2012), 341), 'the assumption needed for financial markets to have a real effect [on firms' choices] via the transmission of information is not that real decision makers are less informed than traders, but only that they do not have perfect information about every decision-relevant factor, and so outsiders may possess some *incremental* information that is useful to them. Thus, real decision makers may be the most informed agents in the economy about the firm, but there are still aspects about which they can learn from outsiders.'

Our empirical results consistently show that the contribution of R&D to the market value of firms is positively affected by the SEZs program, which suggests that the program achieved its aim to increase incentives for innovation.

In addition to an overall positive effect of the SEZs program on the market value of R&D, we find that R&D intensive firms, defined as the firms with the highest level of R&D investment scaled by their total assets, benefit most from the program. This finding suggests that, although the SEZs program reached its aim to promote innovation, it also increased the gap between less and more R&D-intensive firms. This is an unintended consequence of the policy with potential long-term implications for competition, innovation, and growth.

We contribute to the literature in several ways. First, while previous studies on China focused mostly on the national institutional context (e.g. P. H. Hsu, Tian, and Xu 2014; X. Liu and White 2001; Motohashi and Yun 2007), the SEZs program allows us to focus on the regional institutional environment. Regional variation in China's institutional context has only recently started to gain attention (e.g. Barbieri, Pollio, and Prota 2020; Barca, Mccann, and Rodriguez-Pose 2012; Howell 2019; Hua, Partridge, and Sun 2023; Koster et al. 2019; Y. Liu, Wu, and Zhu 2023; Lu, Wang, and Zhu 2019; Ma, Cao, and Zhao 2023) and it highlights the importance of analysing the *local* institutional environment where firms operate as a key determinant of their innovation activities (Barasa et al. 2017; Rodríguez-Pose and Zhang 2020) rather than the national ones.

Second, we specifically contribute to the literature on China's SEZs program which has been shown to positively affect regional productivity (Howell 2019; Ma, Cao, and Zhao 2023), firm productivity and performance (Koster et al. 2019; Y. Liu, Wu, and Zhu 2023), foreign direct investment (FDI), output growth (Barbieri, Pollio, and Prota 2020; Lu, Wang, and Zhu 2019; J. Wang 2013; Zheng et al. 2015) and the number of patent applications, patent grants, and patent citations (Wu, Liu, and Huang 2021). Other than Wu, Liu, and Huang (2021), we do not focus directly on innovative inputs and outputs, such as R&D and patents, but show how the contribution of R&D to the market value of firms changes if firms are involved in the SEZs program. A change in the contribution of R&D to the market value of firms reflects a change in firms' incentives to innovate. This approach provides further insights because the innovation process, from the search for financial means and human talent to the launch of a final product, is long-term, so that the immediate change in R&D investment or output only captures a small part of the actual effect, while changes to firm's incentive can provide insights into expected long-term benefits.

Third, we provide nuanced evidence on the effects of the SEZs program and show that its positive effect on the contribution of R&D to the market value of firms is greater for R&D intensive firms in the affected regions. While fostering innovation among less R&D intensive firms promotes economic diversity and improves the competitive landscape as these firms often bring fresh ideas and disruptive technologies to the market, driving overall economic growth and dynamism (Gilbert and Newbery 1982; Schumpeter 1942), stimulating R&D intensive firms' innovation activity can lead to a lower overall innovation output because R&D intensive firms have less incentives to innovate as new products would cannibalise profits from existing ones (Arrow 1962; Reinganum 1983). Hence, we show an unintended side effect of the SEZs program in the form of an increasing gap in the incentives for more and less R&D intensive firms to invest in R&D.

Lastly, we contribute to the literature on the market value of R&D by adding evidence for China as a developing economy. Since the majority of studies have focused on developed countries (e.g. Bloom and Van Reenen 2002; Czarnitzki, Hussinger, and Leten 2020; Griliches 1981; Hall, Jaffe, and Trajtenberg 2005; Pindado, De Queiroz, and De La Torre 2010; Toivanen, Stoneman, and Bosworth 2002; see; Czarnitzki, Hall, and Oriani 2006, for a survey), together with Chadha and Oriani (2010) and Kanwar and Hall (2015) who focus on India, we are the first to study the contribution of R&D to the market value of firms for a developing economy. Differently from these two studies, we draw conclusions from a within-country analysis, exploiting sub-regional variation in institutional development through the SEZs program in China.

2. Institutions and the market value of R&D

Institutional contexts directly affect firms' engagement in innovation, and their ability to translate innovative efforts into performance advances (Acharya and Subramanian 2009; Acharya, Baghai, and Subramanian 2013; Donges, Meier, and Silva 2023; Moser 2005; North 1991). Well-developed institutional contexts are characterised by strong and reliable financial and legal systems (Acharya and Subramanian 2009; Acharya, Baghai, and Subramanian 2013) which mitigate the uncertainty surrounding innovation activities and ensure the enforcement of intellectual property rights (IPRs) (Moser 2005) and property rights in general which are crucial for the appropriation of returns (He and Tian 2020).

A well-developed institutional context is characterised by strong enforcement of property rights, including IPRs, a functioning system of laws, rules, and contract enforcement, has efficient intermediary and financial institutions, reliable access to information, access to capital, developed capital markets, and efficient capital allocation systems, well-developed fiscal, monetary, and financial policies, developed banking and regulatory system (see, for example, Fernández and Tamayo 2017; P. H. Hsu, Tian, and Xu 2014; Kumar, Rajan, and Zingales 1999).

The degree of financial development affects the value of innovative assets in several ways (Fernández and Tamayo 2017; Levine 1997). First, more developed equity markets provide firms with greater access to external financing for innovation (P. H. Hsu, Tian, and Xu 2014). Considering the difficulties intrinsic to the valuation of innovative assets, developed equity markets allow information about such assets to reach investors more rapidly providing timely and accurate securities' prices and assets' values. Second, a more liberalised stock market improves innovation via the relaxation of financial constraints, enhanced risk-sharing, and improved corporate governance (Moshirian et al. 2021), which then results, together with higher participation in the marketplace, in a more precise and unbiased value of firms' assets. Third, more efficient capital allocation systems and better access to capital allow firms to engage more strongly in innovation activities (Kumar, Rajan, and Zingales 1999; Xin, Zhang, and Zheng 2017), while whenever access to credit is restricted, large, incumbent firms are in a favourable position as compared to small entrants which are often unable to acquire the required capital for investing in R&D (Davis and Henrekson 1999).

Legal institutions also affect the market value of innovative assets in several ways. First, in regions where IPR protection is only marginally enforced, there is a greater risk that innovative firms might not be able to appropriate the economic rents of their innovation activities (Qian et al. 2017). Considering the direct link between the appropriation of economic rents and innovative assets' value, it is straightforward that a less efficient IPR protection system can lead to a lower market value of innovative assets (Chadha and Oriani 2010; Teece 1986). Second, more enforced property rights facilitate markets for innovation (Spulber 2013), while a low degree of investor protection facilitates information asymmetries and when paired with a weak level of contract enforcement results again in a higher risk of rent misappropriation and lower market value of innovation (Himmelberg, Hubbard, and Love 2004). Third, more developed patent laws have been shown to increase innovation activities in a wide variety of sectors (Moser 2005) and are critical for firms' ability to protect the economic rents derived from their innovation efforts (Cohen, Nelson, and Walsh 2000; Levin et al. 1987).

For all the above, in a well-developed institutional context, the private returns to R&D are expected to be greater as the financing of innovation and the appropriability of its returns are easier. This should be reflected in the contribution of R&D to the market value of firms. In other words, if the institutional context improves, firms should receive a larger contribution of the same level of R&D to their market value as the private returns from R&D can be expected to be greater and more certain.

Hypothesis (H1). The contribution of R&D to the market value of a firm increases after it is included in the SEZs program relative to firms located in regions untargeted by the SEZs program.

3. Special economic zones: the case of China

Launched in 1979 in the cities of Shenzhen, Zhuhai, Shantou, and Xiamen, China's SEZs program, a key industrial policy program aiming at strengthening the performance of China's businesses, was expanded greatly over time and across the country. In fact, since the above-mentioned four SEZs were considered a great success, in 1984, the State Council approved 14 more SEZs in the coastal cities of Qinhuangdao,

Tianjin, and Ningbo (Wu, Liu, and Huang 2021). After 1984, especially after Deng Xiaoping's famous southern tour in 1992, the number of SEZs increased gradually. As of 2018, even though there are considerable debates on the effectiveness of such policies (Barca, Mccann, and Rodriguez-Pose 2012), China counted 2,543 SEZs fairly divided across the 31 different provinces, cities, and independent regions (see Figure A1, Panel A).

SEZs' main objectives are to foster agglomeration economies by building industrial clusters, increasing employment, and attracting technologically advanced industrial facilities by increasing foreign and domestic investments, promoting international trade, and stimulating technological cooperation and innovation in a specific geographic region (Alder, Shao, and Zilibotti 2016). The vast majority of the SEZs (2,039 out of the 2,543 as of 2018) are provincial SEZs (see Figure A1, Panel C), which means that they are under the control of the provincial government which enjoys a certain degree of independence and the freedom to define its own regional policies (Lu, Wang, and Zhu 2019). In fact, each SEZ has its own administration committee and, as their success is directly linked to the political careers of regional government officials, their supervisors are strongly incentivised to strive for the best possible outcome in their respective SEZ (Lu, Wang, and Zhu 2019; Xu 2011). Firms located within a provincial SEZ are granted privileges way beyond the national standard in terms of, e.g. tax deductions and customs duty exemptions, discounted land-use fees, and special treatment in securing bank loans (Lu, Wang, and Zhu 2019). In addition, as China's business environment was characterised by weak institutions, including poor protection of private property rights, limited access to financial resources, and weak infrastructure, which limited entrepreneurship and innovation, provincial SEZs were seen as experiments aimed at providing better institutions, reducing inefficiencies in the institutional environment, and improving innovation (Alder, Shao, and Zilibotti 2016; Lu, Wang, and Zhu 2019). In fact, SEZs provide greatly improved conditions for firms to engage in innovation activities. The SEZs administration committees implemented a variety of policies to encourage innovation by reducing innovation costs, e.g. providing more favourable tax credit policies than the national standard for R&D, and strengthening innovation incentives, e.g. setting up innovation and development funds, venture capital funds, and industrial investment funds to provide financial subsidies to support innovation in enterprises and providing support to the development of financial intermediaries and patent intermediaries, and providing enterprises with training about intellectual property (Wu, Liu, and Huang 2021). All these initiatives are conducive to higher innovation incentives and efficiency as well as an improved business environment for innovation and entrepreneurship.

It is worthwhile noting that next to provincial SEZs, China established national-level SEZs which can grant greater privileges to the resident firms (Alder, Shao, and Zilibotti 2016). As those national-level SEZs are much less frequent and different in nature as the responsible committees enjoy more authority for managing the zones and their aim is to benefit the entire economy, we only focus on provincial SEZs following Lu, Wang, and Zhu (2019). It is important to note that provincial- and state-level SEZs cannot overlap (Lu, Wang, and Zhu 2019).

Another important fact is that, while the first SEZ regions were chosen based on favourable geographical locations, industrial conditions, and human capital (J. Wang 2013), in later years, the program became more representative of the spatial distribution and less subject to selection biases (Lu, Wang, and Zhu 2019). This less selective allocation of SEZs is important for our empirical analysis which would be otherwise affected by selection biases (Lu, Wang, and Zhu 2019). In response, we only focus on SEZs established in later years following Lu, Wang, and Zhu (2019).⁵

4. Methodology, data and variables

4.1. Methodology

4.1.1. The market value model

R&D activities undertaken by firms represent investments yielding intangible assets often referred to as 'knowledge stock' (Bloom 2007). When these assets are recognised to positively affect a firm's expected

⁵We address this potential issue further with our control group of geographically close firms as discussed in the methodology section. We would also like to emphasise that potential regional efforts to improve the local infrastructure in order to qualify for the SEZ program are not reflected in the firm variables of interest for this study (see Figure 4).

6 🛭 😔 K. HUSSINGER AND L. PALLADINI

future cash flows, they should be positively reflected in the firm's market value (Griliches 1981). Hence, a firm's investment in R&D should be reflected in its market value, meaning that not only innovation outputs, such as patents and new products, affect the market value of a firm, but also innovation inputs which turn into profits in the future years (Hall, Jaffe, and Trajtenberg 2005). Following this intuition, Griliches (1981) formulated the so-called 'market value model' which allows quantifying the contribution of innovative assets to the market value of a firm (Hall 1993a). Since the market value itself is a forward-looking measure that incorporates future expected returns, Griliches (1981) approach has the advantage of avoiding timing issues stemming from the time distance of R&D costs and revenues which usually occur years after the investment (Czarnitzki, Hall, and Oriani 2006; Hall 2000). The timing aspect is especially important in the case of innovative assets where the timeline of the innovative process from the idea to the product launch is not known to the corporate researcher and corporate insider (Ahuja, Coff, and Lee 2005).

A further advantage is that Griliches' market value model allows us to empirically identify the value of specific assets separately. Drawing from the hedonic pricing model, the market value model views firms as bundles of assets and capabilities whose value is defined as the present discounted value of the expected future cash flows (Czarnitzki, Hall, and Oriani 2006). The value of a firm (V_{it}) is therefore described as a function of its physical (A_{it}) and knowledge (K_{it}) assets⁶

$$V_{it} = q_t (A_{it} + \gamma K_{it})^{\sigma} \tag{1}$$

Under the assumption of constant returns to scale ($\sigma = 1$), Equation (1) can be written as a market value equation where Q_{it} represents the market value of the firm defined as the assets' value over their replacement costs:

$$\log Q_{it} = \log \frac{V_{it}}{A_{it}} = \log q_t + \log \left(1 + \gamma \frac{K_{it}}{A_{it}}\right)^{\sigma}$$
(2)

Equation (2) can be estimated directly using nonlinear least squares (NLLS) (e.g. Czarnitzki, Hussinger, and Leten 2020). Alternatively, ordinary least squares (OLS) regression models can be applied using a linear approximation of the model if the approximation $log(1 + x) \sim x$ is valid (e.g. Griliches 1981; Jaffe 1986; Cockburn and Griliches 1988; Hall 1993a; 1993b; Czarnitzki et al. 20006; Simeth and Cincera 2016; Hussinger and Pacher 2019). The linear approximation has the advantage to allow for an easy inclusion of firm fixed effects and interaction terms which we need to test our hypothesis. We check the difference between log(1 + x) and x for our innovative assets' variables and find that the approximation holds for our sample.,⁷ Hence, we estimate the following equation:

$$logQ_{it} \approx \gamma \frac{K_{it}}{A_{it}}$$
(3)

To test our hypothesis, we use firms' R&D investment scaled by the assets' book value to capture the knowledge assets K_{it} of a firm.

The market value model developed by Griliches (1981) requires that stock markets are efficient enough to reflect the future expected return of investments. Carpenter, Lu, and Whitelaw (2021) show that, since the reforms of the early 2000s, stock prices in China have become as informative about firm future profits as they are in the U.S. and they have become much more efficient since (Chong, Lam, and Yan 2012). Adding to these general findings, recent evidence shows a positive association between the market value of Chinese firms and patents (D. H. Hsu, Hsu, and Zhao 2021; P.-H. Hsu, Wang, and Wu 2013).

4.1.2. Estimation methods

To estimate the market value model, we use a geographical BD approach and show robustness for a coarsened exact matching and random sampling approach. We always distinguish between 'treated' firms, i.e. those firms located in a geographical region targeted by the SEZs program, and 'untreated' firms located in non-SEZ regions. For the geographical BD approach, we select a control group of untreated firms that are geographically close to regions that were turned into SEZs (Holmes 1998; Keele and Titiunik 2015; Lu, Wang, and Zhu 2019).

 $^{{}^{6}}q_{t}$ is the market valuation coefficient of the firm's total assets; γ is the relative shadow value of knowledge capital to tangible assets.:

⁷As shown in Table A1, the measure for R&D investment over assets in our sample has a mean of 0.019 and a maximum value of 0.1. The approximation is, whenever appropriate, standard practice (see for example Simeth and Cincera 2016).

From the pool of untreated firms, we select firms located between five and seven kilometres⁸ away from the treated firms at the time of the treatment, but are not included in the SEZ region.⁹ The geographical BD approach is based on the assumption that economic conditions alter abruptly at the borders of an SEZ due to differences in policies, regulations, or infrastructure, while at the same time, the firms themselves are very similar on both sides.

Since the SEZs were established in different years, we use stacked linear fixed effects regressions on an unbalanced, stacked panel dataset for the estimation following e.g. Baker, Larcker, and Wang (2022) and Cengiz et al. (2019).

As mentioned above, we employ two alternative methods for which we define different control groups of untreated firms. For the coarsened exact matching (CEM) approach, we define a control group of one 'identical' untreated firm for each treated firm based on observable characteristics, i.e. industry sector, province, and firm size, as of the pre-treatment year. This approach shows the robustness of our results accounting for potential differences in the observable characteristics between treated and control firms. For the random sampling approach, we use 500 randomly selected groups of three untreated firms per treated firm. This approach improves the analysis by providing a more representative estimate of the entire population, as it draws samples from the full pool of untreated firms. This enhances the generalisability of the findings to the broader population.

4.1.3. Specification and estimation

Since SEZs are established in different years, we combine a staggered difference-in-differences (DiD) setup with Griliches' market value model resulting in the following empirical specification:

$$\log Q_{it} = \beta_1 post_{it} + \beta_2 treat_i \times post_{it} + \beta_3 \frac{R\&D_{it}}{\text{total assets}_{it}} + \beta_4 \frac{R\&D_{it}}{\text{total assets}_{it}} \times post_{it} + \beta_5 \frac{R\&D_{it}}{\text{total assets}_{it}} \times treat_i + \beta_6 \frac{R\&D_{it}}{\text{total assets}_{it}} \times treat_i \times post_{it} + X_{it} + \delta_t + \lambda_i + \varepsilon_{it}$$

$$(4)$$

which includes time (δ_t) and firm (λ_i) fixed effects as well as a dummy *post*_{it} which takes a value of 1 after a treated firm is in a SEZ and 0 otherwise; the respective control firm takes the same values in the same years. In addition, we include a dummy *treat*_i which takes value 1 for firms located in a SEZ and 0 otherwise, and a set of control variables (X_{it}) . Note that the dummy *treat*_i is time-invariant and, as such, is included in the firm-specific fixed effects λ_i .

Since we are interested in a change in the contribution of R&D to the market value of the firm due to the treatment, we interact the treatment effect, $treat_i \times post_{it}$, with our measure for R&D investment ($R & D_{it}$ /*Total Assets*_{it}). The estimated coefficient shows if and how the contribution of R&D to the firm's market value changed after an SEZ was established relative to the control group.

All our results are based on linear regressions with firm- and year-level fixed effects. Driscoll-Kraay standard errors are used for all estimations. These standard errors have the advantage of (1) being heteroscedasticity consistent, (2) producing estimates robust to very general forms of cross-sectional and temporal dependence, and (3) having significantly better small sample properties than alternative options (Hoechle 2007).

4.2. Data

Yearly data for publicly traded Chinese firms on both Shanghai and Shenzhen Stock Exchanges is gathered from the China Stock Market and Accounting Research (CSMAR) database, one of the leading financial and accounting databases on Chinese firms, and includes R&D investment, total assets, shares distribution, and firms' headquarters location, on top of many other financial and accounting indicators.

The dataset is supplemented with information on SEZs from the China Association of Development Zones which provides information on each SEZ's date of establishment, size, and whether it is a national or provincial-level SEZ.¹⁰ To determine the location of the SEZs we use the

⁸In Table A4 in the Appendix, we report the same results when the control group includes firms between (1) 5 and 10 kilometres, (2) 5 and 15 kilometres, (3) 0.5 and 4 kilometres, and (4) 0.5 and 7 kilometres away from the treated firms.

⁹The control group can even include firms from the same administrative district since SEZs do not necessarily overlap with administrative districts (Lu, Wang, and Zhu 2019).

¹⁰https://www.cadz.org.cn/index.php/Develop/index.html.

Google Maps API. Once we have the geographical coordinates of the SEZs and their sizes, we are able to determine which firms are located in a region that becomes an SEZ during our time window.

Of our initial sample, which comprises the complete set of listed firms in the Chinese market from 1999 to 2021, 483 firms (out of 4,981) became part of an SEZ. Since in China firms are required to disclose their R&D investment only since 2007, CSMAR is unable to provide information on R&D at the firm level before 2007. 384 treated firms remain in our sample. In addition, since the first SEZs were established in regions characterised by favourable geographical locations, industrial conditions, and human capital (J. Wang 2013), and only in later years the program tended to be more representative of the spatial distribution and less subject to selection biases (Lu, Wang, and Zhu 2019), we restrict our analysis to SEZs established between 2008 and 2020. This approach restricts the number of treated firms to 226. In order to conduct meaningful DiD analyses, we only include firms observed at least once before and after the establishment of an SEZ in their location so that we can compare the same firm before and after.¹¹ This leaves us with 179 treated firms in our period of observation. After discarding the top 1% percentile of the main variable of interest ($R c D_{it}$ /*Total Assets_{it}* <0.103), 177 treated firms are located in 49 different SEZs (see Figure 1 below for the firms' geographical locations).¹²

The final sample size varies due to the different control group's selection strategies. For the random sampling, we chose three control firms per treated firm leading to a sample size of 708 firms (note that this random draw is repeated 500 times). For the CEM control group, we chose one control firm per treated firm. Lastly, for the BD design, i.e. our main sample of analysis, we find 305 firms in the neighbouring regions of 109 treated firms. In terms of observations, our final main sample for the BD analysis consists of 836 treated and 7,877 control firm-year observations between 2008 and 2021.¹³



Figure 1. Main analysis – treated and control firms' geographical locations (BD sample).

¹¹Note that the time window is between 8 years before and after the SEZ establishment across all our analysis. Our results are robust to different time windows.

¹²Of these firms, 11 relocated into an SEZ after its establishment. Our results hold when these 11 firms and their control firms are excluded from the sample. The results are available upon request.

¹³Note that some control firms are associated with more than one treated firm.

4.3. Variables

Our dependent variable is the natural logarithm of Tobin's Q defined as the ratio of the firm's market value, i.e. the total number of shares outstanding multiplied by the price per share, over the book value of the firm's tangible assets (Dowell, Hart, and Yeung 2000; Morck and Yeung 1991). Tobin's Q is a forward-looking measure that incorporates future expected profits, i.e. the aggregate market expectations on the returns to current investments (e.g. Griliches 1981; Hall 2000; Hall, Jaffe, and Trajtenberg 2005).

The key independent variable of interest, R&D investment, is measured as the sum of both capitalised and expensed R&D divided by the book value of the assets (Hall, Jaffe, and Trajtenberg 2005; Pindado, De Queiroz, and De La Torre 2010).

A set of control variables is used, including total assets, operating income scaled by total assets, total liabilities over total assets,¹⁴ percentage of state-owned shares, and the number of shares held by the top ten shareholders. These variables are chosen to control for potential firm-specific differences such as firm size, financial health and risks, performances, and corporate governance, among others, and to avoid omitted variables' bias.

To account for the regional institutional environment in which the firm is located, we also include the five main dimensions of the NERI index of marketisation (X. Wang, Fan, and Zhu 2007) as regressors. The NERI index is divided into five broad fields (government and market relations, development of the nonstate enterprise sectors, development of the commodity market, credit market development, and development of legal environment) composed of a total of 23 basic indicators. The index takes values between zero and ten where ten corresponds to the best score. We use the ranks of the provinces for our analysis since the scores increase over time. Ranks take values between one and thirty-one, i.e. the total number of provinces, cities, and independent regions, with thirty-one being the region with the lowest level of institutional development. Table A2 in the Appendix provides the variable definitions. Lastly, we control for time effects using a set of time dummies. Industry dummies are absorbed by the firm-specific effects.

5. Results

5.1. Descriptive statistics

Table 1 shows descriptive statistics for our main sample of treated and their geographically close control firms. Tabel A1 in the Appendix shows descriptive statistics for the full sample which are quite comparable.

Small, insignificant differences in the means suggest that the treated and control firms are similar. Figure 2 shows t-tests for the mean differences between the treated group and the control group which are insignificant.

The mean R&D investment over total assets is about 0.02 for both groups.¹⁵ The average percentage of state-owned shares is about 4.9% with a maximum of about 87.2%. Controlling for state ownership is especially relevant in the case of China because these firms have been shown to differ systematically from other publicly listed firms (e.g. Phi et al. 2021; Shleifer and Vishny 1994; Tong, Junarsin, and Li 2015).

We show the correlation matrix in Table 2 and the industry frequency by treated and control groups in Figure 3.

5.2. Empirical results

5.2.1. Parallel trends

Although we are not interested in the direct effect of the treatment on the dependent variable, i.e. the market value of firms, we conduct a parallel trends analysis to address the concerns that (1) there could have been regional efforts in terms of infrastructure investment to qualify for the SEZs program which would be reflected in the variables of interest at the firm level and (2) idiosyncratic changes in Tobin's Q or R&D investment due to the establishment of the SEZs which might drive our results.

¹⁴More specifically, we consider the book value of interest-bearing liabilities over book value of tangible assets.

¹⁵Figure A2 shows the distribution of *R*&*D_{it}*/*Total Assets_{it}* by firm size where large firms and small firms are defined as above and below the sample mean of the book value of their total assets.

				Treated	duorb							Control	group			
Variables	Mean	SD	Min	25 perc	Median	75 perc	90 perc	Max	Mean	SD	Min	25 perc	Median	75 perc	90 perc	Max
Log Tobin's Q	0.073	0.957	-3.572	-0.53	0.108	0.752	1.248	2.536	-0.085	1.042	-4.554	-0.729	0.026	0.652	1.124	3.457
R&D	0.022	0.018	0	0.009	0.019	0.031	0.046	0.099	0.019	0.019	0	0.005	0.014	0.027	0.043	0.101
investment/Total Assets																
Liabilities/Assets	0.199	0.172	0	0.041	0.175	0.311	0.421	1.056	0.217	0.183	0	0.061	0.182	0.33	0.481	1.05
Income/Total Assets	0.036	0.076	-0.68	0.018	0.038	0.066	0.105	0.271	0.034	0.072	-0.677	0.013	0.031	0.062	0.096	0.384
% of State-Owned Shares	0.043	0.13	0	0	0	0	0.136	0.75	0.05	0.14	0	0	0	0.001	0.141	0.872
% Shares of Top 10 Shareholders	58.443	15.267	15.617	46.769	59.074	70.819	76.918	91.454	58.795	16.588	11.2	46.594	59.329	71.636	79.797	97.499
NERI Index 1 (rank)	6.923	6.409	-	2	4	11	17	31	7.596	5.443	-	2	6	11	14	28
NERI Index 2 (rank)	10.129	6.066	-	4	11	13	19	31	13.895	5.55	-	11	13	19	20	31
NERI Index 3 (rank)	11.327	7.48	-	9	11	14	22	31	18.257	9.069	-	11	19	28	30	30
NERI Index 4 (rank)	7.135	6.633	-	2	S	10	18	31	3.697	5.065	-	-	2	£	11	30
NERI Index 5 (rank)	7.457	6.415	-	m	5	11	18	30	4.954	4.357	-	2	4	5	10	30
Post Dummy	0.544	0.498	0	0	-	-	-	-	0.467	0.499	0	0	0	-	-	-
Total Assets	1199.063	2991.917	22.079	137.066	316.323	808.712	2149.244	27075.017	7209.667	21643.625	13.446	207.159	504.976	2112.726	20087.61	290001.4
R&D investment	18.578	63.947	0	1.987	4.611	10.83	31.872	829.409	75.163	267.462	0.001	2.204	6.189	19.117	103.9	3992.738
R&D investment and total assets are	in units of 1	0,000,000 C	NY. Total I	number of	firms: 414.	Number of	treated firm	109. Numk	ser of contro	l firms: 305. T	Fotal num	oer of obser	vations: 8,7	713. Numbe	r of observa	tions in the

sample).
(BD
statistics
Descriptive
Table

treated group: 836. Number of observations in the control group: 7,877.



Figure 2. T-tests for main variables: treated versus control firms (BD sample).

Figure 4 shows that there is a parallel trend of the dependent variable, the market value of the firms, for the treated and control groups before the establishment of the SEZs. Also, after the establishment of a SEZ, the trend is similar for treated and control firms and not significantly different from each other. While in a standard DiD setting, where we are interested in a change in the dependent variable due to an event, we would expect a significant deviation of the trends after the treatment, an insignificant effect is helpful in our specific context because it suggests that an increase in the contribution of R&D to the market value of the firm due to the SEZs program is not driven by an overall increase of the market value of the firm after the establishment of the SEZ.

Figure 4 further shows that the R&D over assets variable shows a parallel trend for the treated and control groups before and after the establishment of a SEZ. This is also important for our analysis since it suggests that the contribution of R&D to the market value of the firm in the post-SEZ period is not impacted by an increase (or decrease) in R&D investment.

Lastly, this analysis shows that potential regional efforts to qualify for an SEZ program are not reflected in our firm level key variables. This in line with Lu, Wang, and Zhu (2019) who argues that while the first SEZs were subject to selection, the later SEZ regions, which define our sample, were rather randomly selected.¹⁶

5.2.2. Main results

Table 3 presents the main results for the BD approach.

Model (1) shows a lean specification which only includes the post-SEZ dummy and its interaction with the treatment dummy. The treatment dummy is time-invariant hence it is included in the firm specific fixed effects λ_i . All specifications include firm fixed effects as well as time fixed effects. The results show a positive market value effect for the treated firms after the SEZ was established.

Model (2) adds the measures for $R & D_{it}^{17}$ as well as the control variables. As expected, we find a positive, significant effect of $R \mathcal{E} D_{it}$ on Tobin's Q.

Model (3) provides the first test of our hypothesis. The triple interaction term of $R \mathscr{O}_{D_i}$, treat_i, and post_{it} measures the difference in how R&D_{it} contributed to the market value of treated firms as opposed to control firms after the SEZ is established. The positive, significant coefficient suggests that after the SEZ's establishment, R&D_{it} contributed more to the market value of a firm located within its boundaries. Considering that

¹⁶Table A3 in the Appendix provides further suggestive evidence supporting the parallel trend assumption.

¹⁷We refer to our main variable of interest (*R&D_{it}* /*Total* Assets_{it}) simply as *R&D_{it}*

Table 2. Correlation table (BD sample).

	Log	R&D investment/Total	Liabilities/	Income/Total	% of State-Owned	% Shares of Top 10	NERI	NERI	NERI	NERI	NERI	Total
	Tobin's Q	Assets	Assets	Assets	Shares	Shareholders	Index 1	Index 2	Index 3	Index 4	Index 5	Assets
R&D	0.34****											
investment/Total Assets												
Liabilities/Assets	-0.39****	-0.40****										
Income/Total Assets	0.20****	0.17****	-0.28****									
% of State-Owned Shares	-0.33***	-0.06****	0.05****	0.01								
% Shares of Top 10	-0.36***	-0.06****	0.04****	0.15****	0.29****							
Shareholders												
NERI Index 1	-0.02	0.07****	0.06****	0.01	0.08****	0.01						
NERI Index 2	-0.04***	0.08****	0.01	-0.02	0.02	0.03*	0.60****					
NERI Index 3	-0.08****	0.08****	0.03**	0.03**	0.11****	0.07****	0.53****	0.70****				
NERI Index 4	0.06****	-0.06****	0.02	0.02	0.00	-0.07****	0.34***	-0.20****	-0.36***			
NERI Index 5	0.02*	-0.03**	0.05****	-0.02	0.00	-0.07****	0.50****	0.07****	-0.20****	0.79****		
Total Assets	-0.52****	-0.15****	0.13****	-0.05****	0.05****	0.31****	0.10****	0.18****	0.18***	-0.14***	-0.06****	
R&D investment	-0.39****	0.02*	0.04***	-0.03*	0.01	0.26****	0.11****	0.17****	0.16****	-0.12****	-0.05****	0.77****
Signif. codes:***: 0.001, ***	*: 0.01, **: 0.0	15, *: 0.1.										



Figure 3. Industry frequency (BD sample).

Figure 4 shows no increase in $R \notin D_{it}$ for treated firms after the SEZ, this effect is not driven by a potential direct effect of the SEZ on the level of R&D investments. Model (4) shows support for our hypothesis when our list of control variables is included.¹⁸

Since we use a triple interaction term, the interpretation of the model is not straightforward because the estimated treatment effect on the treated depends on R&D investment over assets which occurs several times in our model specification. To illustrate the size of the effect, we plot the linear prediction of the model against $R & D_{it}$ for the treated and the control groups before and after the establishment of the SEZs. Figure 5 shows that, for the treated group, we see a large increase in the predicted contribution of $R & D_{it}$ to the market value of the firms from before to after the establishment of the SEZ only for larger levels of $R & D_{it}$ indicating that the effect of the SEZ is stronger for R&D intensive firms. For less R&D intensive firms, the difference appears to be negligible.

Model (5) shows that our results are robust if a different clustering method for the standard errors is used. Considering that some treated firms are associated with disproportionally more control firms given their geographical proximity to a larger number of firms, we cluster the standard errors at the group level.¹⁹ In Model (6) and (7), we provide further evidence that our results are not driven by the above-mentioned imbalances. In Model (6), we allow each treated firm to have a maximum of three control firms. To do so, we randomly pick three control firms in the pool of firms that are geographically close to the treated firm. In Model (7), we repeat this exercise and allow a maximum of five control firms per treated firm. As all the models report stable and significant coefficient for the triple interaction term of $R & D_{it}$, treat_i, and post_{it}, we interpret this as evidence that SEZs increase the market valuation of the R&D investments of the firms located in the SEZs territories therefore providing incentives for the firms to invest more in innovation.²⁰, ²¹

¹⁸The results are robust for different sets of control variables.

¹⁹When we cluster the standard errors at the firm level, results also hold.

²⁰Table A4 in the Appendix shows that our results are robust when control firms are selected in an area between (1) 5 and 10 kilometres, (2) 5 and 15 kilometres, (3) 0.5 and 4 kilometres, and (4) 0.5 and 7 kilometres away from the treated firms.

²¹Due to the unbalanced nature of our panel dataset Table A5 in the Appendix shows several robustness tests in which we include (1) only firms that are present in the dataset for at least three years, (2) only those firms that are observed always between and including periods -2 and +2, and (3) only in the manufacturing sector, which hosts the largest R&D spenders. Results hold.



Figure 4. Parallel trends – dependent variable and main variable of interest. Notes: The parallel trends are obtained by regressing the two variables of interest (log Tobin's Q and R&D investment over total assets) on the interaction of *treat_i* and a set of dummy variables indicating the temporal distance to the SEZ establishment year. The reference period is t = -1. The same set of control variables as in the main regressions is used.

5.2.3. Placebo test

We run a placebo check in which we move the SEZ establishment date three years earlier. As expected, there is no effect of the placebo treatment on the market value of $R \notin D_{it}$ (Table 4).

5.2.4. Robustness test - coarsened exact matching

Table 5 presents the results for a matched sample. We match 113 treated firms to a comparable control firm using coarsened exact matching (CEM) in the pre-treatment year. Matching criteria include the 3-digit industry code, the province, and three size classes defined along the firms' total assets distribution.²² Our results, as shown in Table 5, are qualitatively unchanged and provide further support for our hypothesis.

5.2.5. Robustness test – random sampling

Table 6 shows robustness for a random sampling approach. We draw 500 random samples of three control firms for each treated firm and run 500 regressions. Standard errors are bootstrapped.

5.2.6. Robustness test – subsidiaries

A potential issue of our analysis is that we rely on headquarter information and do not take into account whether a firm has a subsidiary, a joint venture, or any other type of associated enterprise in a SEZ.²³ To overcome this issue, we extracted the list of subsidiaries from CSMAR and geolocated them using the Google Maps API. Overall, we find 5,119 subsidiaries (out of 27,140) which are located in a SEZ and have a business scope related to R&D. For this robustness check, we added the firms with subsidiaries in a SEZ to our treatment group which leads to a larger treatment group of 606 publicly listed firms that have either a subsidiary or their headquarters in a SEZ.²⁴

²²The small number of treated firms does not allow conducting an exact matching for more matching criteria.

²³Hereafter, for clarity, we refer to all subsidiaries, joint ventures, and any other type of enterprises closely associated with the listed firms as 'subsidiaries'.

²⁴A caveat of the CSMAR data is that the subsidiaries' exact locations are not traced over time. We manually checked whether our subsidiaries changed city in the past and found that 78.3% of the subsidiaries did not change city over time.

Dependent Var.:					Log Tobin's Q		
Model:	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Sample and cluster:	(full sample)	(full sample)	(full sample)	(full sample)	(full sample – clustered s.e. at group- ID level)	(max 3 controls)	(max 5 controls)
Variables	coef. (s.e.)	coef. (s.e.)	coef. (s.e.)	coef. (s.e.)	coef. (s.e.)	coef. (s.e.)	coef. (s.e.)
Post	0.0646*** (0.0148)	0.0478*** (0.0138)	0.0494* (0.0266)	0.0542 (0.0323)	0.0542*** (0.0201)	0.1574** (0.0524)	0.1278** (0.0504)
Treat x Post	0.0844**	0.1219***	-0.0971	-0.0371	-0.0371 (0.1027)	-0.1354	-0.1031
R&D/Assets	(0.0257)	4.911***	2.857***	4.772***	4.772***	7.794***	7.195***
R&D/Assets x Treat		(1.037)	0.5475	-1.218	-1.218	-3.823	-3.706
R&D/Assets x Post			0.7796	-0.3524	-0.3524	-1.442 (1.794)	-1.616
R&D/Assets x Treat			7.723**	7.164**	7.164*	7.280**	7.671**
Controls Fixed-Effects:	No	Yes	No	Yes	Yes	Yes	Yes
Firm Year	Yes Yes	Yes Yes	Yes Yes	Yes Yes	Yes Yes	Yes Yes	Yes Yes
Fit Statistics S.E. type	Drisco	Drisco	DrisKra.	DrisKra.	Group-ID	DrisKra.	DrisKra.
	Kra. (L = 1)	Kra. (L = 1)	(L = 1)	(L = 1)		(L = 1)	(L = 1)
Observations R ²	8,713 0.78773	8,713 0.83602	8,713 0.78909	8,713 0.83629	8,713 0.83629	2,507 0.77754	2,976 0.80268
Within R ²	0.00283	0.22969	0.00923	0.23094	0.23094	0.23132	0.22016

Table 3. Main results from a BD approach.

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms: 109. Number of control firms: 305. Results are based on OLS estimations with yearand firm-level fixed effects. Driscoll-Kraay standard-errors are shown in parentheses. *Full sample* indicates we use *all* the control firms located in the neighbouring regions (between 5 and 7 kilometres) of the treated firms. In Models 6 and 7, out of all the control firms associated with each treated firm, we randomly select up to 3 and 5 control firms, respectively. Model 5 shows robustness of our results for different clustering of the standard errors. As some treated firms are associated to a disproportionally larger number of control firms, we cluster the standard errors at the overall group level (a group is identified as the treated firm and all associated control firms).

Dependent Var.:		Log To	bin's Q	
Model	(1)	(2)	(3)	(4)
	coef.	coef.	coef.	coef.
Variables	(s.e.)	(s.e.)	(s.e.)	(s.e.)
Post	-0.0114	-0.0107	-0.0326	-0.0055
	(0.0213)	(0.0178)	(0.0404)	(0.0359)
Treat x Post	0.0837	0.0773	0.0366	0.0458
	(0.0618)	(0.0455)	(0.1171)	(0.0921)
R&D/Assets		5.462***	3.178**	4.865***
		(1.288)	(1.418)	(1.485)
R&D/Assets x Treat			7.948**	5.071*
			(3.051)	(2.358)
R&D/Assets x Post			1.186	-0.2266
			(1.276)	(1.244)
R&D/Assets x Treat x Post			1.518	1.132
			(3.015)	(2.349)
Controls	No	Yes	No	Yes
Fixed-Effects:				
Firm	Yes	Yes	Yes	Yes
Year	Yes	Yes	Yes	Yes
Fit Statistics				
S.E. type	DrisKra. $(L = 1)$	Driscoll-Kra. ($L = 1$)	DrisKra. $(L = 1)$	Driscoll-Kra. (L = 1)
Observations	7,924	7,924	7,924	7,924
R ²	0.79768	0.84384	0.79933	0.84402
Within R ²	0.00041	0.22846	0.00854	0.22934

Table 4. Placebo test based on the BD.

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms: 109. Number of control firms: 305. The placebo treatment is set to happen three years before the actual treatment. Results are based on OLS estimations with year- and firm-level fixed effects. Driscoll-Kraay standard-errors are shown in parentheses.



Figure 5. Linear predictions.

Dependent Var.:		Log Tobi	n's Q	
Model:	(1)	(2)	(3)	(4)
	coef.	coef.	coef.	coef.
Variables	(s.e.)	(s.e.)	(s.e.)	(s.e.)
Post	0.0013	-0.0563	0.1012	0.0371
	(0.0584)	(0.0491)	(0.0836)	(0.0754)
Treat x Post	0.1753***	0.2054***	-0.0804	0.0046
	(0.0364)	(0.0263)	(0.0869)	(0.0419)
R&D/Assets		7.444***	11.61***	8.103***
		(1.187)	(2.646)	(2.188)
R&D/Assets x Treat			-5.666	-0.6483
			(3.502)	(2.457)
R&D/Assets x Post			-5.173**	-3.957*
			(1.957)	(2.155)
R&D/Assets x Treat x Post			11.92***	8.396***
			(3.139)	(1.810)
Controls	No	Yes	No	Yes
Fixed-Effects:				
Firm	Yes	Yes	Yes	Yes
Year	Yes	Yes	Yes	Yes
Fit Statistics				
S.E. type	DriscoKra. $(L = 1)$	DriscoKra. (L = 1)	DrisKra. (L = 1)	DrisKra. (L = 1)
Observations	1,983	1,978	1,983	1,978
R ²	0.69571	0.76387	0.70463	0.76545
Within R ²	0.00843	0.22906	0.03751	0.23420

Tal	ble	5.	Rol	bustness	test -	 coarsened 	exact	matc	hin	ıg
-----	-----	----	-----	----------	--------	-------------------------------	-------	------	-----	----

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms: 113. Number of control firms: 113. Matching criteria: province, industry sector, three firm size classes (as measured by total assets), and year. Results are based on OLS estimations with year- and firm-level fixed effects. Driscoll-Kraay standard-errors are presented in parentheses.

Table 7 shows the results for the CEM approach and random control samples.²⁵ They are in line with the main results.

²⁵Note that a robustness check for the subsidiary sample based on BD regressions would violate the key assumption of the BD approach which demands that the treated and control firms are located closely to the geographical border. We only have consolidated firm information, i.e. we do not have information on the level of the subsidiary. Since our subsidiary sample of treated firms includes subsidiaries located in a SEZ with a headquarter located far away, the main assumption of the BD design is, hence, violated.

Tal	ole 6	5. Ro	obustness	test –	500	random	control	groups.
-----	-------	-------	-----------	--------	-----	--------	---------	---------

Dependent Var.:		Log 1	obin's Q	
Model:	(1)	(2)	(3)	(4)
	coef.	coef.	coef.	coef.
Variables	(s.e.)	(s.e.)	(s.e.)	(s.e.)
Post	0.001	0.001	0.001	0.001
	(0.003)	(0.003)	(0.003)	(0.003)
Treat x Post	0.120***	0.134***	-0.032	0.030
	(0.027)	(0.023)	(0.029)	(0.031)
R&D/Assets		6.158***	17.849	7.956
		(1.260)	(14.215)	(11.959)
R&D/Assets x Treat			-9.128	-2.146
			(14.291)	(12.136)
R&D/Assets x Post			-8.890***	-5.175***
			(0.523)	(0.540)
R&D/Assets x Treat x Post			10.040***	6.521***
			(0.384)	(0.546)
Controls	No	Yes	No	Yes
Fixed-Effects:				
Firm	Yes	Yes	Yes	Yes
Year	Yes	Yes	Yes	Yes

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms: 177. Number of control firms: 531. Standard errors are based on 500 random draws for the control group. Results are based on stacked OLS estimations with year-and firm-level fixed effects.

Table 7. Ro	bustness	test - su	bsidia	ries
-------------	----------	-----------	--------	------

Dependent Var.:				Log Tol	oin's Q			
Model	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Sample	(CEM)	(CEM)	(CEM)	(CEM)	(Random)	(Random)	(Random)	(Random)
Variables	coef.	coef.	coef.	coef.	coef.	coef.	coef.	coef.
	(s.e.)	(s.e.)	(s.e.)	(s.e.)	(s.e.)	(s.e.)	(s.e.)	(s.e.)
Post	0.0785**	0.0853**	0.0953**	0.1541***	-0.005	-0.004	-0.003	-0.000
	(0.0277)	(0.0298)	(0.0372)	(0.0467)	(0.004)	(0.004)	(0.007)	(0.022)
Treat x Post	-0.0003	-0.0142	-0.1017	-0.1473**	0.094***	0.092***	0.008	0.014
	(0.0312)	(0.0338)	(0.0622)	(0.0678)	(0.023)	(0.019)	(0.026)	(0.022)
R&D/Assets		6.410***	10.32***	10.17***		3.279***	18.873**	8.689*
		(0.5566)	(1.421)	(1.833)		(0.390)	(9.424)	(4.795)
R&D/Assets x Treat			-6.645***	-6.924**			-14.476	-7.859
			(2.171)	(2.652)			(9.482)	(4.894)
R&D/Assets x Post			-1.212	-3.249*			-0.192	-0.237
			(1.202)	(1.814)			(0.281)	(0.274)
R&D/Assets x Treat			4.582**	6.014**			3.447***	3.596***
x Post			(2.132)	(2.506)			(0.361)	(0.299)
Controls	No	Yes	No	Yes	No	Yes	No	Yes
Fixed-Effects:								
Firm	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Fit Statistics								
S.E. type	DrisKra.	DriscKra.	DrisKra.	DriscKra.				
	(L = 1)	(L = 1)	(L = 1)	(L = 1)				
Observations	6,092	6,067	6,092	6,067				
R ²	0.68662	0.73012	0.69349	0.73102				
Within R ²	0.00248	0.13562	0.02436	0.13851				

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms in the CEM sample: 355. Number of control firms in the CEM sample: 355. Matching criteria: province, industry sector, three firm size classes (as measured by total assets), and year. Number of treated firms in the repeated random sampling: 606. Number of control firms in the repeated random sampling: 1818. Standard errors for the random sample are based on 500 random draws for the control group. Results are based on stacked OLS estimations with year- and firm-level fixed effects.

5.2.7. Heterogeneity analysis

We further attempt to disentangle the effect of the different types of SEZs. To do so, we divide the SEZs in our sample into tech and non-tech related SEZs. More precisely, we identify all provincial-level SEZ defined as 'High-tech Industrial Park', 'High-tech Industrial Development Zone', 'Science and Technology Park', 'Science and Technology Innovation Industry Functional Zone', and 'High-tech Zone' as those with the main goal of fostering domestic firms' innovation. We find that 45% of our treated firms (49) are located

Dependent Var.:	Log To	bin's Q
Model:	(1)	(2)
Sample:	Tech	No tech
	coef.	coef.
Variables	(s.e.)	(s.e.)
R&D/Assets	4.744*	3.659*
	(2.510)	(1.812)
Post	0.0887**	0.0140
	(0.0354)	(0.0331)
R&D/Assets x Treat	3.071	-0.3838
	(3.651)	(5.743)
R&D/Assets x Post	-0.4819	-0.2346
	(1.540)	(1.557)
Treat x Post	-0.1559	0.1654
	(0.0885)	(0.1577)
R&D/Assets x Treat x Post	5.952**	3.180
	(2.394)	(5.906)
Controls	Yes	Yes
Fixed-Effects:		
Firm	Yes	Yes
Year	Yes	Yes
Fit Statistics		
S.E. type	DrisKra.	DriKra.
	(L = 1)	(L = 1)
Observations	3,515	4,991
R ²	0.80267	0.82509
Within R ²	0.20047	0.19265

Table 8. Heterogeneity analysis: tech SEZs versus non-tech SEZs.

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms (Model 1): 49. Number of control firms (Model 1): 133. Number of treated firms (Model 2): 60. Number of control firms (Model 2): 172. Results are based on OLS estimations with year- and firm-level fixed effects. Driscoll-Kraay standard-errors are presented in parentheses.

within this SEZs group. We then run subsample analyses as reported in Table 8 below. As expected, the main effect is found in the subsample of tech-related SEZs (Model (1)); the main effect vanishes for firms located in non-high-related SEZs (Model (2)).

Furthermore, throughout the paper we argue that the higher level of institutional development generated by the SEZ drives our main results. If this is true, then we should observe an increase in the market valuation of R&D investment after the SEZ establishment especially in those regions that, ex-ante, i.e. before the SEZ establishment, show the lowest level of legal and financial institutional development. To test this, we further split the sample according to the mean of the NERI index before the SEZ establishment. We use the overall NERI index (Models (1) and (2) of Table 9) as well as the NERI subindices that measure the level of the development of the credit market (Models (3) and (4) of Table 9) and of the legal institutions (Models (5) and (6) of Table 9).

We find the triple interaction term to be significant only for firms located in SEZs established in provinces that before their establishment experienced the lowest levels of institutional development (Model (1)). We further find that SEZs are especially helpful in contexts where the credit market development is weak (Model (3)). Differently from the credit market development, we find that for any level of legal environment development, the establishment of an SEZ improves the firms' incentives to innovate (Models (5) and (6)). A possible explanation is that this occurs due to the overall low level of legal institutional development throughout China's national territory.

A last heterogeneity check concerns R&D intensive and less R&D intensive firms. We run a subsample analysis where we split the sample of firms around the mean of R&D intensity ($R \notin D_{it}$ /Total Assets_{it}). We find that R&D intensive firms are the ones benefiting the most from the SEZs (Model (7)) providing further evidence of the possible unintended consequences of the SEZ program which might widen the gap between less and more R&D intensive innovators, potentially impacting competition and long-term growth.

 Table 9. Heterogeneity analysis: institutional development and R&D intensity.

Dependent Var.:	Log Tobin's Q							
Model:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Sample	Low pre-SEZ institutional development	High pre-SEZ institutional development	Low pre-SEZ credit market institutional development	High pre-SEZ credit market institutional development	Low pre-SEZ legal institutional development	High pre-SEZ legal institutional development	R&D intensive firms	Less R&D intensive firms
Variables	coef.	coef.	coef.	coef.	coef.	coef.	coef. (s.e.)	coef.
	0.2385	5 306***	4 5 2 2	A 705***	4 007*	5 268***	2 026	15 02***
Post	(3.581) 0.0353	(1.201) 0.0464	(3.404) 0.0362	(0.9670) 0.0517	(2.313) -0.0107	(1.143) 0.0676*	(1.979) -0.0369	(2.617) 0.0477***
R&D/Assets	(0.0809)	(0.0286)	(0.0768) -2.634	(0.0325)	(0.0489)	(0.0360)	(0.0461)	(0.0148) _17.09***
x Treat	(6 915)	(3 869)	(6 215)	(3 776)	(5 118)	(2 212)	(3.825)	(4 566)
R&D/Assets	0.9026	-0.0140	0.1490	-0.2473	-0.2434	0.0838	2.865**	-2.247
x Post	(2.189)	(1.095)	(2.311)	(1.183)	(1.408)	(1.201)	(1.179)	(2.167)
Treat x Post	-0.1371	0.1179	-0.1348	0.1508	-0.2214***	0.2391**	-0.1222	-0.0150
	(0.1192)	(0.1323)	(0.1482)	(0.1533)	(0.0672)	(0.0968)	(0.1332)	(0.1085)
R&D/Assets	10.89**	2.660	11.48**	0.9018	7.375**	4.301*	8.185**	6.833
x Treat x Post	(4.582)	(3.782)	(5.000)	(4.404)	(3.000)	(2.196)	(3.374)	(6.887)
Controls Fixed-Effects:	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Fit Statistics								
S.E. type	DrisKra.	DrisKra.	DrisKra.	DriscKra.	DriscoKra.	DrisKra.	DrisKra.	DriscKra.
	(L = 1)	(L = 1)	(L = 1)	(L = 1)	(L = 1)	(L = 1)	(L = 1)	(L = 1)
Observations	1,644	6,765	1,954	6,582	3,149	5,260	3,803	4,753
R [∠]	0.78911	0.83112	0.78067	0.85180	0.81142	0.82922	0.73596	0.82599
Within R ²	0.23165	0.19157	0.23235	0.24101	0.21255	0.18765	0.21160	0.20476

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms overall: 109. Number of control firms overall: 305. Results are based on OLS estimations with year- and firm-level fixed effects. Driscoll-Kraay standard-errors are presented in parentheses.

6. Discussion

One of the main objectives of the SEZs program in China is to foster innovation in certain geographical regions by improving those regions' institutional environment (Alder, Shao, and Zilibotti 2016). We investigate whether the SEZs program in China succeeded in its objective to incentivise innovation by analysing the contribution of R&D to the market value of the firms located in the targeted geographical regions.

Our difference-in-differences geographical boundary discontinuity results from a stacked sample around the establishment of the SEZs in different regions in China show that, when institutions become stronger through the SEZs program, the contribution of R&D to the market value of firms increases. This finding suggests that the SEZs program was successful at stimulating innovation. These results are robust when different estimation frameworks are used and towards several robustness checks.

Our results carry considerable implications for industrial policies in China as they establish a direct link between the SEZs program and firms' incentives to innovate. Differently from other studies that focus on the innovation inputs and outputs, i.e. R&D investment and patents, derived by such programs, we argue that analysing the change in the contribution of R&D to the firm's value is a valuable contribution because the market value is a forward-looking measure which is able to capture incentives to innovate rather than innovation input or innovation success.

One caveat that our analysis reveals, however, is that the effect of the SEZs program is greater for R&D intensive firms and weaker for firms with a smaller R&D intensity located in SEZ regions. This raises questions about the long-term effects of the SEZs program. Fostering innovation among small firms promotes economic diversity and improves the competitive landscape as small innovators often bring fresh ideas and disruptive technologies to the market, driving overall economic growth and dynamism (Gilbert and Newbery 1982; Schumpeter 1942). If, however, R&D intensive firms' innovation activity is stimulated by the SEZs program, the overall effect on innovation can be much smaller than what it could

20 🔶 K. HUSSINGER AND L. PALLADINI

have been because R&D intensive firms have fewer incentives to innovate as new products cannibalise profits from existing ones (Arrow 1962; Reinganum 1983). In order to evaluate the long-term effect of the SEZs program, an analysis of the specific measures taken in each SEZ, i.e. the specific policy mix, would be required. Due to the limited number of publicly listed firms per SEZ, this would be rather a topic for a case study on the regional level.

Furthermore, the SEZs program aimed to reach its goals of business growth and innovation by providing better institutions and by reducing inefficiencies in the institutional environment (Alder, Shao, and Zilibotti 2016; Lu, Wang, and Zhu 2019). By showing that firms located within SEZs observe a larger contribution of their R&D efforts to their market value, we add to the literature on the market value of firms in developing economies and how more developed institutions can increase the market assessment of innovative assets by decreasing institutional uncertainties. The majority of studies in this field have focused on developed countries (see Czarnitzki, Hall, and Oriani 2006, for a survey), while we, together with Chadha and Oriani (2010) and Kanwar and Hall (2015) who focus on India, are the first to study the contribution of R&D to the market value of firms for a developing economy. Our study further benefits from at least two methodological advantages: (1) we draw conclusions from a within-country analysis, exploiting sub-regional variation in institutional development through the SEZs program in China, and (2) we draw conclusions based on arguably exogenous shocks (see the arguments above related to the exogeneity of the SEZ program implementation especially in later years).

A further limitation of our study is that China is an economy that experiences a variety of national, provincial, and local policies, the implementation of which typically varies across different geographical regions. This makes it virtually impossible to control for all of these changes in a regression analysis. Our empirical approach, however, to some extent, handles confounding factors. First, the BD design alleviates the concern of confounding policy changes by focusing on units near the regional border where the SEZ is implemented. What is crucial in our setting is that the SEZ borders do not coincide with administrative borders (Lu, Wang, and Zhu 2019) so it is unlikely that a treated and control firm are not affected equally by potentially confounding factors to some extent by leveraging variation in the timing of implementation and across regions. These approaches minimise the influence of confounding national events because such events do not align with each treated region's establishment date of an SEZ. Since all these approaches, although relying on completely different assumptions, lead to very similar results, we are confident that our results are reflective of an actual effect of SEZs.

Data availability statement

The data used for this study is available from the China Stock Market and Accounting Research (CSMAR): https://data. csmar.com/. CSMAR is a commercial data provider and the authors do not have permission to share the data.

Disclosure statement

No potential conflict of interest was reported by the author(s).

References

- Acharya, V. V., R. P. Baghai, and K. V. Subramanian. 2013. "Labor Laws and Innovation." *The Journal of Law and Economics* 56 (4): 997–1037. https://doi.org/10.1086/674106.
- Acharya, V. V., and K. V. Subramanian. 2009. "Bankruptcy Codes and Innovation." *The Review of Financial Studies* 22 (12): 4949–4988. https://doi.org/10.1093/rfs/hhp019.
- Ahuja, G., R. W. Coff, and P. M. Lee. 2005. "Managerial Foresight and Attempted Rent Appropriation: Insider Trading on Knowledge of Imminent Breakthroughs." *Strategic Management Journal* 26 (9): 791–808. https://doi.org/10.1002/ smj.474.

Alder, S., L. Shao, and F. Zilibotti. 2016. "Economic Reforms and Industrial Policy in a Panel of Chinese Cities." *Journal of Economic Growth* 21 (4): 305–349. https://doi.org/10.1007/s10887-016-9131-x.

- Arora, A., A. Fosfuri, and A. Gambardella. 2001. *Markets for Technology: The Economics of Innovation and Corporate Strategy*. Cambridge MA: The MIT Press.
- Arrow, K. J. 1962. "Economic Welfare and the Allocation of Resources for Invention." In *The Rate and Direction of Inventive Activity*, edited by R. Nelson, 609–625. Princeton: National Bureau of Economic Research and Princeton University Press.
- Bai, J., T. Philippon, and A. Savov. 2016. "Have Financial Markets Become More Informative?" Journal of Financial Economics 122 (3): 625–654. https://doi.org/10.1016/j.jfineco.2016.08.005.
- Baker, A. C., D. F. Larcker, and C. C. Wang. 2022. "How Much Should We Trust Staggered Difference-In-Differences Estimates?" *Journal of Financial Economics* 144 (2): 370–395. https://doi.org/10.1016/j.jfineco.2022.01.004.
- Barasa, L., J. Knoben, P. Vermeulen, P. Kimuyu, and B. Kinyanjui. 2017. "Institutions, Resources and Innovation in East Africa: A Firm Level Approach." *Research Policy* 46 (1): 280–291. https://doi.org/10.1016/j.respol.2016.11.008.
- Barbieri, E., C. Pollio, and F. Prota. 2020. "The Impacts of Spatially Targeted Programmes: Evidence from Guangdong." *Regional Studies* 54 (3): 415–428. https://doi.org/10.1080/00343404.2019.1635688.
- Barca, F., P. Mccann, and A. Rodriguez-Pose. 2012. "The Case for Regional Development Intervention: Place-Based versus Place-Neutral Approaches." *Journal of Regional Science* 52 (1): 134–152. https://doi.org/10.1111/j.1467-9787. 2011.00756.x.
- Bloom, N. 2007. "Uncertainty and the Dynamics of R&D." American Economic Review 97 (2): 250–255. https://doi.org/ 10.1257/aer.97.2.250.
- Bloom, N., and J. Van Reenen. 2002. "Patents, Real Options and Firm Performance." *The Economic Journal* 112 (478): C97–C116. https://doi.org/10.1111/1468-0297.00022.
- Bond, P., A. Edmans, and I. Goldstein. 2012. "The Real Effects of Financial Markets." Annual Review of Financial Economics 4 (1): 339–360. https://doi.org/10.1146/annurev-financial-110311-101826.
- Bond, P., and I. Goldstein. 2015. "Government Intervention and Information Aggregation by Prices." *Journal of Finance* 70 (6): 2777–2812. https://doi.org/10.1111/jofi.12303.
- Carpenter, J. N., F. Lu, and R. F. Whitelaw. 2021. "The Real Value of China's Stock Market." *Journal of Financial Economics* 139 (3): 679–696. https://doi.org/10.1016/j.jfineco.2020.08.012.
- Cengiz, D., A. Dube, A. Lindner, and B. Zipperer. 2019. "The Effect of Minimum Wages on Low-Wage Jobs." *The Quarterly Journal of Economics* 134 (3): 1405–1454. https://doi.org/10.1093/qje/qjz014.
- Chadha, A., and R. Oriani. 2010. "R&D Market Value Under Weak Intellectual Property Rights Protection: The Case of India." *Scientometrics* 82 (1): 59–74. https://doi.org/10.1007/s11192-009-0042-x.
- Chong, T. T. L., T. H. Lam, and I. K. M. Yan. 2012. "Is the Chinese Stock Market Really Inefficient?" *China Economic Review* 23 (1): 122–137. https://doi.org/10.1016/j.chieco.2011.08.003.
- Cockburn, I., and Z. Griliches. 1988. "Industry Effects and Appropriability Measures in the Stock market's Valuation of R&D and Patents." *American Economic Review* 78 (2): 419–423. https://www.jstor.org/stable/1818161.
- Cohen, W. M., R. R. Nelson, and J. P. Walsh. 2000. "Protecting Their Intellectual Assets: Appropriability Conditions and Why US Manufacturing Firms Patent (Or Not)." *NBER Working Paper*, w7552.
- Czarnitzki, D., B. H. Hall, and R. Oriani. 2006. "The Market Valuation of Knowledge Assets in US and European Firms." In *The management of intellectual property. New Horizons in intellectual property series*, 111–131. Cheltenham, England: Edward Elgar Publishing. https://www.e-elgar.com/shop/usd/the-management-of-intellectual-property-9781845421120.html/.
- Czarnitzki, D., K. Hussinger, and B. Leten. 2020. "How Valuable are Patent Blocking Strategies?" *Review of Industrial Organization* 56 (3): 409–434. https://doi.org/10.1007/s11151-019-09710-9.
- Dang, T. V., and Z. Xu. 2018. "Market Sentiment and Innovation Activities." *Journal of Financial and Quantitative Analysis* 53 (3): 1135–1161. https://doi.org/10.1017/S0022109018000078.
- Davis, S. J., and M. Henrekson. 1999. "Explaining National Differences in the Size and Industry Distribution Employment." Small Business Economics 12 (1): 59-83. https://doi.org/10.1023/A:1008078130748.
- Donges, A., J. M. Meier, and R. C. Silva. 2023. "The Impact of Institutions on Innovation." *Management Science* 69 (4): 1951–1974. https://doi.org/10.1287/mnsc.2022.4403.
- Dowell, G., S. Hart, and B. Yeung. 2000. "Do Corporate Global Environmental Standards Create or Destroy Market Value?" *Management Science* 46 (8): 1059–1074. https://doi.org/10.1287/mnsc.46.8.1059.12030.
- Duranton, G., and A. J. Venables. 2018. *Place-Based Policies for Development* (No. w24562). Cambridge, MA: National Bureau of Economic Research.
- Farole, T., and G. Akinci, Eds. 2011. Special Economic Zones: Progress, Emerging Challenges, and Future Directions. Washington, DC: World Bank Publications.
- Fernández, A., and C. E. Tamayo. 2017. "From Institutions to Financial Development and Growth: What are the Links?" *Journal of Economic Surveys* 31 (1): 17–57. https://doi.org/10.1111/joes.12132.
- Gilbert, R. J., and D. M. Newbery. 1982. "Preemptive Patenting and the Persistence of Monopoly." *The American Economic Review* 72 (3): 514–526.
- Griliches, Z. 1981. "Market Value, R&D, and Patents." *Economics Letters* 7 (2): 183–187. https://doi.org/10.1016/0165-1765(87)90114-5.
- Hall, B. H. 1993a. "The Value of Intangible Corporate Assets: An Empirical Study of the Components of Tobin's Q." Berkeley, CA: University of California at Berkeley Dept. of Economics Working Paper No. 93-207.

- Hall, B. H. 1993b. "The stock market's valuation of R&D investment during the 1980's." *The American Economic Review* 83 (2): 259-264.
- Hall, B. H. 2000. "Innovation and Market Value." In *Productivity, Innovation and Economic Performance*, edited by R. Barrell, G. Mason, and M. O'Mahoney, 177–199. Cambridge: Cambridge University Press.
- Hall, B. H., A. Jaffe, and M. Trajtenberg. 2005. "Market Value and Patent Citations." *The RAND Journal of Economics* 36 (1): 16–38. http://www.jstor.org/stable/1593752.
- Hartwell, C. A. 2018. "Bringing the Benefits of David to Goliath: Special Economic Zones and Institutional Improvement." *Regional Studies* 52 (10): 1309–1321. https://doi.org/10.1080/00343404.2017.1346371.
- He, J., and X. Tian. 2020. "Institutions and Innovation." Annual Review of Financial Economics 12 (1): 377–398. https://doi.org/10.1146/annurev-financial-032820-083433.
- Himmelberg, C. P., R. G. Hubbard, and I. Love. 2004. "Investor Protection, Ownership, and the Cost of Capital." World Bank Policy Research Working Paper, No. 2834.
- Hoechle, D. 2007. "Robust Standard Errors for Panel Regressions with Cross-Sectional Dependence." *The Stata Journal* 7 (3): 281–312. https://doi.org/10.1177/1536867X0700700301.
- Holmes, T. J. 1998. "The Effect of State Policies on the Location of Manufacturing: Evidence from State Borders." *Journal of Political Economy* 106 (4): 667–705. https://doi.org/10.1086/250026.
- Hottenrott, H., B. H. Hall, and D. Czarnitzki. 2016. "Patents as Quality Signals? The Implications for Financing Constraints on R&D." *Economics of Innovation & New Technology* 25 (3): 197–217. https://doi.org/10.1080/ 10438599.2015.1076200.
- Howell, A. 2019. "Heterogeneous Impacts of China's Economic and Development Zone Program." *Journal of Regional Science* 59 (5): 797–818. https://doi.org/10.1111/jors.12465.
- Hsu, D. H., P.-H. Hsu, and Q. Zhao. 2021. "Rich on Paper? Chinese firms' Academic Publications, Patents, and Market Value." *Research Policy* 50 (9): 104319. https://doi.org/10.1016/j.respol.2021.104319.
- Hsu, P. H., X. Tian, and Y. Xu. 2014. "Financial Development and Innovation: Cross-Country Evidence." *Journal of Financial Economics* 112 (1): 116–135. https://doi.org/10.1016/j.jfineco.2013.12.002.
- Hsu, P.-H., C. Wang, and C. Wu. 2013. "Banking Systems, Innovations, Intellectual Property Protections, and Financial Markets: Evidence from China." *Journal of Business Research* 66 (12): 2390–2396. https://doi.org/10.1016/j.jbusres. 2013.05.025.
- Hua, Y., M. Partridge, and W. Sun. 2023. "Pollution Effects of Place-Based Policy: Evidence from China's Development-Zone Program." *Journal of Regional Science* 63 (3): 703–727. https://doi.org/10.1111/jors.12637.
- Hussinger, K., and S. Pacher. 2019. "Information Ambiguity, Patents and the Market Value of Innovative Assets." *Research Policy* 48 (3): 665–675. https://doi.org/10.1016/j.respol.2018.10.022.
- Jaffe, A. B. 1986. "Technological Opportunity and Spillovers of R&D: Evidence from firms' Patents, Profits, and Market Value." *The American Economic Review* 76 (5): 984–1001. http://www.jstor.org/stable/1816464.
- Kanwar, S., and B. H. Hall. 2015. "The Market Value of R&D in Weak Innovation Regimes: Evidence from India." *NBER Working Paper*, w21196.
- Keele, L. J., and R. Titiunik. 2015. "Geographic Boundaries as Regression Discontinuities." *Political Analysis* 23 (1): 127–155. https://doi.org/10.1093/pan/mpu014.
- Koster, H. R., F. F. Cheng, M. Gerritse, and F. G. van Oort. 2019. "Place-Based Policies, Firm Productivity, and Displacement Effects: Evidence from Shenzhen, China." *Journal of Regional Science* 59 (2): 187–213. https://doi.org/ 10.1111/jors.12415.
- Kumar, K. B., R. G. Rajan, and L. Zingales. 1999. "What Determines Firm Size?" NBER Working Paper, w7208.
- Lev, B., B. Sarath, and T. Sougiannis. 2005. "R&D Reporting Biases and Their Consequences." Contemporary Accounting Research 22 (4): 977–1026. https://doi.org/10.1506/7XMH-QQ74-L6GG-CJRXquery.
- Levin, R. C., A. K. Klevorick, R. R. Nelson, S. G. Winter, R. Gilbert, and Z. Griliches. 1987. "Appropriating the Returns from Industrial Research and Development." *Brookings Papers on Economic Activity* 1987 (3): 783–831.
- Levine, R. 1997. "Financial Development and Economic Growth: Views and Agenda." *Journal of Economic Literature* 35 (2): 688–726.
- Levitas, E., and M. A. McFadyen. 2009. "Managing Liquidity in Research-Intensive Firms: Signaling and Cash Flow Effects of Patents and Alliance Activities." *Strategic Management Journal* 30 (6): 659–678. https://doi.org/10.1002/smj.762.
- Liu, X., and S. White. 2001. "Comparing Innovation Systems: A Framework and Application to China's Transitional Context." *Research Policy* 30 (7): 1091–1114. https://doi.org/10.1016/S0048-7333(00)00132-3.
- Liu, Y., and Y. Jin. 2022. "Special Economic Zones, Export Status, and firms' Productivity: Theory and Evidence from China." *Review of Development Economics* 26 (3): 1338–1360. https://doi.org/10.1111/rode.12883.
- Liu, Y., Y. Wu, and X. Zhu. 2023. "Development Zones and firms' Performance: The Impact of Development Zones on firms' Performance for a Chinese Industrial Cluster." *Regional Studies* 57 (5): 868–879. https://doi.org/10.1080/ 00343404.2022.2107192.
- Long, C. 2002. Patent Signals, 625-679. Chicago: The University of Chicago Law Review.
- Lu, Y., J. Wang, and L. Zhu. 2019. "Place-Based Policies, Creation, and Agglomeration Economies: Evidence from China's Economic Zone Program." American Economic Journal, Economic Policy 11 (3): 325–360. https://doi.org/10. 1257/pol.20160272.

- Ma, T., X. Cao, and H. Zhao. 2023. "Development Zone Policy and High-Quality Economic Growth: Quasi-Natural Experimental Evidence from China." *Regional Studies* 57 (3): 590–605. https://doi.org/10.1080/00343404.2022. 2093342.
- Morck, R., and B. Yeung. 1991. "Why Investors Value Multinationality." *Journal of Business* 64 (2): 165–187. https://doi.org/10.1086/296532.
- Moser, P. 2005. "How Do Patent Laws Influence Innovation? Evidence from Nineteenth-Century world's Fairs." *American Economic Review* 95 (4): 1214–1236. https://doi.org/10.1257/0002828054825501.
- Moshirian, F., X. Tian, B. Zhang, and W. Zhang. 2021. "Stock market liberalization and innovation." *Journal of Financial Economics* 139 (3): 985–1014. https://doi.org/10.1016/j.jfineco.2020.08.018.
- Motohashi, K., and X. Yun. 2007. "China's Innovation System Reform and Growing Industry and Science Linkages." *Research Policy* 36 (8): 1251–1260. https://doi.org/10.1016/j.respol.2007.02.023.
- Munari, F., and R. Oriani. 2005. "Privatization and Economic Returns to R & D Investments." *Industrial and Corporate Change* 14 (1): 61–91. https://doi.org/10.1093/icc/dth049.
- North, D. C. 1991. "Institutions." Journal of Economic Perspectives 5 (1): 97-112. https://doi.org/10.1257/jep.5.1.97.
- Phi, N. T. M., F. Taghizadeh-Hesary, C. A. Tu, N. Yoshino, and C. J. Kim. 2021. "Performance Differential Between Private and State-Owned Enterprises: An Analysis of Profitability and Solvency." *Emerging Markets Finance and Trade* 57 (14): 3913–3928. https://doi.org/10.1080/1540496X.2020.1809375.
- Pindado, J., V. De Queiroz, and C. De La Torre. 2010. "How Do Firm Characteristics Influence the Relationship Between R&D and Firm Value?" *Financial Management* 39 (2): 757–782. https://doi.org/10.1111/j.1755-053X.2010. 01091.x.
- Qian, C., H. Wang, X. Geng, and Y. Yu. 2017. "Rent Appropriation of Knowledge-Based Assets and Firm Performance When Institutions are Weak: A Study of Chinese Publicly Listed Firms." *Strategic Management Journal* 38 (4): 892–911. https://doi.org/10.1002/smj.2522.
- Reinganum, J. F. 1983. "Uncertain Innovation and the Persistence of Monopoly." *The American Economic Review* 73 (4):741–748. http://www.jstor.org/stable/1816571.
- Rodríguez-Pose, A., and M. Zhang. 2020. "The Cost of Weak Institutions for Innovation in China." *Technological Forecasting and Social Change* 153:119937. https://doi.org/10.1016/j.techfore.2020.119937.
- Schumpeter, J. A. 1942. Capitalism, Socialism, and Democracy, 132-145. Vol. 36. New York: Harper & Row.
- Shleifer, A., and R. W. Vishny. 1994. "Politicians and Firms." The Quarterly Journal of Economics 109 (4): 995–1025. https://doi.org/10.2307/2118354.
- Simeth, M., and M. Cincera. 2016. "Corporate Science, Innovation, and Firm Value." *Management Science* 62 (7): 1970–1981. https://doi.org/10.1287/mnsc.2015.2220.
- Song, Y., R. Deng, R. Liu, and Q. Peng. 2020. "Effects of Special Economic Zones on FDI in Emerging Economies: Does Institutional Quality Matter?" Sustainability 12 (20): 8409. https://doi.org/10.3390/su12208409.
- Spulber, D. F. 2013. "How Do Competitive Pressures Affect Incentives to Innovate When There is a Market for Inventions?" *Journal of Political Economy* 121 (6): 1007–1054. https://doi.org/10.1086/674134.
- Teece, D. J. 1986. "Profiting from Technological Innovation: Implications for Integration, Collaboration, Licensing and Public Policy." *Research Policy* 15 (6): 285–305. https://doi.org/10.1016/0048-7333(86)90027-2.
- Toivanen, O., P. Stoneman, and D. Bosworth. 2002. "Innovation and the Market Value of UK Firms, 1989–1995." Oxford Bulletin of Economics and Statistics 64 (1): 39–61. https://doi.org/10.1111/1468-0084.00002.
- Tong, S., E. Junarsin, and C. Li. 2015. "A Comparative Study of Chinese SOE Firm's Boards and Private Firm's Boards." Annals of Economics & Finance 16 (2): 291-314.
- United Nations Conference on Trade and Development. 2019. World Investment Report 2019: Special Economic Zones. UN.
- Wang, J. 2013. "The Economic Impact of Special Economic Zones: Evidence from Chinese Municipalities." *Journal of Development Economics* 101:133–147. https://doi.org/10.1016/j.jdeveco.2012.10.009.
- Wang, X., G. Fan, and H. Zhu. 2007. "Marketisation in China, Progress and Contribution to Growth." In *China: Linking Markets for Growth*, edited by R. Garnaut, and L. Song, 30–44. Canberra, Australia: Australian National University E Press.
- Wu, M., C. Liu, and J. Huang. 2021. "The Special Economic Zones and Innovation: Evidence from China." China Economic Quarterly International 1 (4): 319–330. https://doi.org/10.1016/j.ceqi.2021.11.004.
- Xin, F., J. Zhang, and W. Zheng. 2017. "Does Credit Market Impede Innovation? Based on the Banking Structure Analysis." *International Review of Economics & Finance* 52:268–288. https://doi.org/10.1016/j.iref.2017.01.014.
- Xu, C. 2011. "The Fundamental Institutions of China's Reforms and Development." *Journal of Economic Literature* 49 (4): 1076–1151. https://doi.org/10.1257/jel.49.4.1076.
- Zheng, S., W. Sun, J. Wu, and M. E. Kahn. 2015. The Birth of Edge Cities in China: Measuring the Spillover Effects of Industrial Parks (Working Paper No. w21378). Cambridge, MA: National Bureau of Economic Research (NBER).

Appendix

Table A1 below provides the descriptive statistics for the full sample included in the boundary discontinuity analysis. Table A2 below provides a description of all the variables used in the paper.

Table A3 provides further evidence in support of the parallel trend assumption. Models (1) and (2) show the market value model only for the years before and only for the years after the SEZ, respectively. The interaction term of $R & D_{it}$ and *treat_i* indicates whether the market values R&D differently between treated and control groups before (Model (1)) and after the treatment (Model (2)). Model (1) shows an insignificant coefficient of the interaction terms suggesting that there is no difference in the contribution of R&D to the market value of treated firms and firms in the control group before the treatment. Model (2) shows a significant and positive difference in the years after the treatment indicating that the contribution of R&D to the market value of treated firms is larger compared to control firms after the SEZ.

Model (3) shows the market value regressions for the treated group only, while Model (4) shows the regressions only for the control group. The coefficient of interest is the interaction of $R & D_{it}$ and $post_{it}$. Model (3) shows that the establishment of the SEZs changes the contribution of $R & D_{it}$ to the market values of the firm only for the treated firms, while there is no effect for the control firms (Model (4)). We interpret this as further suggestive evidence that the parallel trend assumption holds.

In Table A4 below, we run several robustness checks where we select the control group of geographically proxime firms between (1) 5 and 10 kilometres, (2) 5 and 15 kilometres, 26 (3) 0.5 and 4 kilometres, and (4) 0.5 and 7 kilometres. The results are robust.

Tab	le A1.	Descriptive	statistics	(full	BD	samp	ble)
-----	--------	-------------	------------	-------	----	------	-----	---

Variables	Mean	SD	Min	25 perc	Median	75 perc	90 perc	Мах
		50		25 pere		75 pere	Jo pere	
Log Tobin's Q	-0.07	1.035	-4.554	-0.715	0.032	0.666	1.13	3.457
R&D investment/Total Assets	0.019	0.019	0	0.005	0.015	0.027	0.043	0.101
Liabilities/Assets	0.215	0.182	0	0.06	0.181	0.328	0.476	1.056
Income/Total Assets	0.034	0.072	-0.68	0.014	0.032	0.062	0.097	0.384
% of State-Owned Shares	0.049	0.139	0	0	0	0	0.141	0.872
% Shares of Top 10 Shareholders	58.761	16.465	11.2	46.594	59.283	71.562	79.241	97.499
NERI Index 1 (rank)	7.532	5.547	1	2	9	11	15	31
NERI Index 2 (rank)	13.533	5.71	1	11	13	19	20	31
NERI Index 3 (rank)	17.592	9.159	1	10	19	28	30	31
NERI Index 4 (rank)	4.027	5.333	1	1	2	3	12	31
NERI Index 5 (rank)	5.194	4.653	1	2	4	5	11	30
Treat Dummy	0.096	0.295	0	0	0	0	0	1
Post Dummy	0.474	0.499	0	0	0	1	1	1
Total Assets	6632.958	20675.75	13.446	198.046	473.494	1825.045	18926.93	290001.4
R&D investment	69.734	255.619	0	2.191	5.974	18.353	90.177	3992.738

R&D investment and total assets are in units of 10,000,000 CNY. Total number of firms: 414. Number of treated firms: 109. Number of control firms: 305. Total number of observations: 8,713. Number of observations in the treated group: 836. Number of observations in the control group: 7,877.

Table A2. Variable definitions	Ta
--------------------------------	----

Variable	Definition
Log Tobin's Q	Number of A shares multiplied by the year-end price per share divided by total assets' book value (log-transformed)
Total assets	Total book value of assets is divided by 10,000,000 CNY
R&D investment	Capitalized and expensed R&D divided by 10,000,000 CNY
State Owned shares	% of state-owned shares
R&D investment/Total	R&D Investment divided by Total Assets
Assets	
Liabilities/Assets	Book value of interest-bearing liabilities over book value of tangible assets
Top 10 shareholders	Shares held by the top 10 shareholders
Income over Assets	Net Income divided by Total Assets
Treat	Dummy taking value of 1 if a firm is located in a SEZ; 0 otherwise
Post	Dummy taking the value of 1 after the treated firm or the corresponding control firm is located in a SEZ; 0 otherwise
NERI Index 1 (rank)	Degree to which a region is regulated by the market and not by the central party (from 1 to 31)
NERI Index 2 (rank)	Degree of development of non-state-owned sector (from 1 to 31)
NERI Index 3 (rank)	Degree of market freedom (for example, whether prices are determined by the market or by the government) (from 1 to 31)
NERI Index 4 (rank)	Degree of development of the credit market (from 1 to 31)
NERI Index 5 (rank)	Degree of development of the legal environment (from 1 to 31)
Industry code	3 digits industry codes grouped into 10 categories

²⁶To avoid strong imbalances in our analysis due to the fact that some treated firms are associated with a large number of control firms, we randomly pick between three and maximum five control firms per treated firm.

Dependent Variable:	Log Tobin's Q							
Model:	(1)	(2)	(3)	(4)				
Subsamples:	Only before	Only after	Only treated	Only control				
	coef.	coef.	coef.	coef.				
Variables	(s.e.)	(s.e.)	(s.e.)	(s.e.)				
R&D/Assets	6.637***	1.868**	3.518	4.558***				
	(0.8256)	(0.8413)	(2.838)	(1.141)				
R&D/Assets x Treat	-2.457	9.722***						
	(3.616)	(1.755)						
Post			0.0683	0.0505				
			(0.0801)	(0.0306)				
R&D/Assets x Post			4.972*	-0.1500				
			(2.663)	(1.228)				
Controls	Yes	Yes	Yes	Yes				
Fixed-Effects:								
Firm	Yes	Yes	Yes	Yes				
Year	Yes	Yes	Yes	Yes				
Fit statistics								
Observations	4,582	4,131	836	7,877				
R ²	0.85957	0.90899	0.80853	0.84071				
Within R ²	0.25479	0.23758	0.34505	0.22544				

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms overall: 109. Number of control firms overall: 305. Results are based on OLS estimations with year- and firm-level fixed effects. Driscoll-Kraay standard-errors are shown in parentheses.

	Table /	A4. BD	analysis	with	different	distance	gaps	(BD	samp	ole)
--	---------	---------------	----------	------	-----------	----------	------	-----	------	------

Dependent Var.:		Log To	bin's Q			
Model:	(1)	(2)	(3)	(4)	(5)	(6)
Sample:	Between 5 and 10	Between 5 and 10	Between 5 and 15	Between 5 and 15	Between 0.5 and 4	Between 0.5 and 7
	kms	kms	kms	kms	kms	kms
	max 3 controls	max 5 controls	max 3 controls	max 5 controls	no max	no max
Variables	coef.	coef.	coef.	coef.	coef.	coef.
	(s.e.)	(s.e.)	(s.e.)	(s.e.)	(s.e.)	(s.e.)
R&D/Assets	6.622***	6.266***	7.284***	5.665***	10.34***	6.564***
	(1.503)	(1.915)	(1.265)	(1.838)	(1.132)	(0.8692)
post	0.1016**	0.0550	0.1011**	0.0266	0.0558***	0.0534**
	(0.0359)	(0.0511)	(0.0445)	(0.0577)	(0.0167)	(0.0232)
R&D/Assets x treat	-0.8131	-0.5657	-1.266	0.5409	-6.459*	-1.951
	(2.954)	(2.916)	(3.317)	(3.150)	(3.156)	(3.260)
R&D/Assets x post	-0.4930	-0.1371	0.0142	0.4583	-1.023	–0.5593
	(1.822)	(1.473)	(1.212)	(1.498)	(0.7864)	(0.7409)
treat x post	-0.0871	—0.0358	-0.0483	-0.0361	0.0912	0.0261
	(0.0632)	(0.0816)	(0.0767)	(0.0717)	(0.0657)	(0.0702)
R&D/Assets x treat	4.251*	4.073*	4.389**	4.236**	6.347** (2.567)	7.065**
Controls Fixed-Effects:	Yes	Yes	Yes	Yes	Yes	Yes
Firm	Yes	Yes	Yes	Yes	Yes	Yes
Year	Yes	Yes	Yes	Yes	Yes	Yes
Fit Statistics S.F. type	DrisKra. (I =1)	Dris-Kra. (I=1)	Dris-Kra. (I=1)	Dris-Kra. (I=1)	Dris-Kra. (I =1)	Dris-Kra. (I=1)
Observations	3,463	4,299	4,027	5,033	6,531	17,655
R ²	0.79747	0.78888	0.76435	0.76215	0.83099	0.83264
Within R ²	0.21684	0.20775	0.21669	0.17202	0.28574	0.22658

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms between 5 and 10 kms: 132. Number of control firms between 5 and 10 kms (max 3 controls): 310. Number of control firms between 5 and 10 kms (max 5 controls): 407. Number of treated firms between 5 and 15 kms: 144. Number of control firms between 5 and 15 kms (max 3 controls): 365. Number of control firms between 5 and 15 kms (max 5 controls): 407. Number of treated firms between 5 and 15 kms (max 3 controls): 365. Number of control firms between 5 and 15 kms (max 5 controls): 494. Number of treated firms between 0.5 and 4 kms: 105. Number of control firms between 0.5 and 4 kms: 245. Number of treated firms between 0.5 and 7 kms: 135. Number of control firms between 0.5 and 7 kms: 496. Results are based on OLS estimations with year- and firm-level fixed effects. Driscoll-Kraay standard-errors are presented in parentheses.

Table A5 shows various specifications addressing the unbalanced nature of the sample. Model (1) shows a robustness test in which we include only firms that are present in the dataset for at least three years. Similarly, Model (2) shows results for only those firms that are observed always between and including periods -2 and +2. In addition, in Model (3), we show robustness of our results for firms only in the manufacturing sector.

26 🛞 K. HUSSINGER AND L. PALLADINI

	Table A5.	Further	robustness	tests:	different	subsam	ples
--	-----------	---------	------------	--------	-----------	--------	------

Dependent Var.:		Log Tobin's Q	
Model:	(1)	(2)	(3)
Sample	(firms observed at least for three years)	(firms observed from period –2 to period +2 around the SEZ establishment)	(manufacturing)
	coef.	coef.	coef.
Variables	(s.e.)	(s.e.)	(s.e.)
R&D/Assets	4.773***	4.617***	8.230***
	(1.098)	(1.143)	(1.008)
Post	0.0544	0.0670	0.0737*
	(0.0323)	(0.0423)	(0.0406)
R&D/Assets x Treat	-1.249	-2.672	-5.239
	(3.340)	(3.575)	(3.651)
R&D/Assets x Post	-0.3642	-2.044	-0.9121
	(1.263)	(1.558)	(1.477)
Treat x Post	-0.0478	-0.2081**	-0.1897*
	(0.0985)	(0.0950)	(0.0968)
R&D/Assets x Treat x Post	7.326**	11.19**	13.63***
	(3.347)	(3.988)	(3.988)
Controls	Yes	Yes	Yes
Fixed effects			
Firm	Yes	Yes	Yes
Year	Yes	Yes	Yes
Fit Statistics			
S.E. type	Driscoll-Kra. (L=1)	Driscoll-Kra. (L=1)	Driscoll-Kra. (L=1)
Observations	8,686	5,852	4,465
R ²	0.83571	0.83360	0.77350
Within R ²	0.23066	0.21968	0.23188

Signif. codes: ***: 0.01, **: 0.05, *: 0.1. Number of treated firms (Model 1): 103. Number of control firms (Model 1): 283. Number of treated firms (Model 2): 44. Number of control firms (Model 2): 178. Number of treated firms (Model 3): 74. Number of control firms (Model 3): 165.Results are based on OLS estimations with year- and firm-level fixed effects. Driscoll-Kraay standard-errors are presented in parentheses.

Figure A1, Panel A, below provides an overview of the location of all the existing SEZs in China. Panels B and C show the spatial distribution of national and provincial SEZs, respectively. Despite the fact that there seems to be an overlap between national and provincial SEZs in the Figure, Lu, Wang, and Zhu (2019),



Figure A1. Geographical distribution of all SEZs.

pp. 329) clarify that "geographically, national and provincial SEZs are mutually exclusive—a location cannot be both a provincial and a national SEZ at the same time".

Figure A2 below provides the distribution of R&D investment, i.e. R&D investment over total assets, by firm size. We define large and small firms by splitting the BD sample at the mean of the book value of the total assets of the firms. Treated and control firms' information is provided on the right and left panels, respectively.



Figure A2. Distribution of R&D investment by firm size (BD sample).